Subset Selection

Nate Wells

Math 243: Stat Learning

October 21st, 2020

Outline

In today's class, we will...

- Investigate algorithms for selecting good subsets of predictors
- Discuss penalized regression as an alternate method of model selection

Section 1

Subset Selection

Nate Wells (Math 243: Stat Learning)

Suppose we wish to find a linear model for Y with p predictors X_1, \ldots, X_p .

Suppose we wish to find a linear model for Y with p predictors X_1, \ldots, X_p .

How do we determine the optimal collection of predictors?

Suppose we wish to find a linear model for Y with p predictors X_1, \ldots, X_p .

How do we determine the optimal collection of predictors?

Determine an appropriate selection criteria.

Suppose we wish to find a linear model for Y with p predictors X_1, \ldots, X_p .

How do we determine the optimal collection of predictors?

Determine an appropriate selection criteria.

• Cross-validation: Computationally expensive

Suppose we wish to find a linear model for Y with p predictors X_1, \ldots, X_p .

How do we determine the optimal collection of predictors?

Determine an appropriate selection criteria.

• Cross-validation: Computationally expensive

Suppose we wish to find a linear model for Y with p predictors X_1, \ldots, X_p .

How do we determine the optimal collection of predictors?

Determine an appropriate selection criteria.

- Cross-validation: Computationally expensive
- Adjusted R^2 : Penalizes non-helpful predictors, but may overestimate test error rate.

Suppose we wish to find a linear model for Y with p predictors X_1, \ldots, X_p .

How do we determine the optimal collection of predictors?

Determine an appropriate selection criteria.

- Cross-validation: Computationally expensive
- Adjusted R^2 : Penalizes non-helpful predictors, but may overestimate test error rate.
- C_p : penalizes training RSS by typical discrepancy between test and training.

$$C_p = \frac{1}{n} (\mathrm{RSS} + 2d\hat{\sigma}^2)$$

Suppose we wish to find a linear model for Y with p predictors X_1, \ldots, X_p .

How do we determine the optimal collection of predictors?

Determine an appropriate selection criteria.

- Cross-validation: Computationally expensive
- Adjusted R^2 : Penalizes non-helpful predictors, but may overestimate test error rate.
- C_p : penalizes training RSS by typical discrepancy between test and training.

$$C_p = rac{1}{n}(\mathrm{RSS} + 2d\hat{\sigma}^2)$$

• Akaike information criterion (AIC): uses method of maximum likelihood, assuming Normal errors

$$AIC = \frac{1}{n\hat{\sigma}^2} (RSS + 2d\hat{\sigma}^2)$$

Suppose we wish to find a linear model for Y with p predictors X_1, \ldots, X_p .

How do we determine the optimal collection of predictors?

Determine an appropriate selection criteria.

- Cross-validation: Computationally expensive
- Adjusted R^2 : Penalizes non-helpful predictors, but may overestimate test error rate.
- C_p : penalizes training RSS by typical discrepancy between test and training.

$$C_p = \frac{1}{n} (\mathrm{RSS} + 2d\hat{\sigma}^2)$$

• Akaike information criterion (AIC): uses method of maximum likelihood, assuming Normal errors

$$AIC = \frac{1}{n\hat{\sigma}^2} (RSS + 2d\hat{\sigma}^2)$$

 Bayesian information criterion (BIC): uses method of maximum likelihood and Bayes' Rule

$$BIC = \frac{1}{n\hat{\sigma}^2} (RSS + \ln nd\hat{\sigma}^2)$$

With p predictors, there are a total of 2^p possible MLR models.

• There are $\binom{p}{k}$ models using exactly k of p predictors

With p predictors, there are a total of 2^p possible MLR models.

• There are $\binom{p}{k}$ models using exactly k of p predictors

Theoretically, we can find the best model by fitting each possible model and selecting the best via appropriate selection criteria (C_p , AIC, BIC, R^2 , CV)

With p predictors, there are a total of 2^p possible MLR models.

• There are $\binom{p}{k}$ models using exactly k of p predictors

Theoretically, we can find the best model by fitting each possible model and selecting the best via appropriate selection criteria (C_p , AIC, BIC, R^2 , CV)

Downsides?

With p predictors, there are a total of 2^p possible MLR models.

• There are $\binom{p}{k}$ models using exactly k of p predictors

Theoretically, we can find the best model by fitting each possible model and selecting the best via appropriate selection criteria (C_p , AIC, BIC, R^2 , CV)

Downsides?

• Computation time and storage grows exponentially in p

With p predictors, there are a total of 2^p possible MLR models.

• There are $\binom{p}{k}$ models using exactly k of p predictors

Theoretically, we can find the best model by fitting each possible model and selecting the best via appropriate selection criteria (C_p , AIC, BIC, R^2 , CV)

Downsides?

- Computation time and storage grows exponentially in p
- May have low marginal improvement despite number of models fitted

We use the regsubsets function in the leaps library.

• regsubsets uses the same syntax as 1m. The summary function outputs the best set of variables for the given number of predictors

- regsubsets uses the same syntax as 1m. The summary function outputs the best set of variables for the given number of predictors
- Be default, regsubsets only returns up to the best eight models. But nvmax can be used to return as many variables as desired

- regsubsets uses the same syntax as 1m. The summary function outputs the best set of variables for the given number of predictors
- Be default, regsubsets only returns up to the best eight models. But nvmax can be used to return as many variables as desired
- Best is determined by *RSS*.

- regsubsets uses the same syntax as 1m. The summary function outputs the best set of variables for the given number of predictors
- Be default, regsubsets only returns up to the best eight models. But nvmax can be used to return as many variables as desired
- Best is determined by RSS.

```
library(palmerpenguins)
library(leaps)
penguins<-penguins %>% drop_na()
```

```
best_subset<-regsubsets(body_mass_g ~. , data = penguins, nvmax = 8)</pre>
```

Summary of regsubsets

• Stars indicate variable is included in model

##	Subset selection object							
##	Call: regsubsets.formula(body_mass_g ~ ., data = penguins, nvmax = 8)							
##	9 Variables (and intercept)							
##	Forced in Forced out							
##	speciesChinstrap	FALSE	FALSE					
##	speciesGentoo	FALSE	FALSE					
##	islandDream	FALSE	FALSE					
##	islandTorgersen	FALSE	FALSE					
##	bill_length_mm	FALSE	FALSE					
##	bill_depth_mm	FALSE	FALSE					
##	flipper_length_mm	FALSE	FALSE					
##	sexmale	FALSE	FALSE					
##	year	FALSE	FALSE					
##	1 subsets of each size up to 8							
##	Selection Algorithm: exhaustive							
##	speciesChins	strap speci	esGentoo islandDream isl	andTorgersen				
##	1 (1) " "							
##	2 (1) " "	"*"						
##	3 (1) " "	"*"						
##	4 (1) " "	"*"						
##	5 (1) "*"	"*"						
##	6 (1) "*"	"*"						
##	7 (1) "*"	"*"						
##	8 (1) "*"	"*"	" " "*"					
##			pth_mm flipper_length_mm					
##			"*"					
##	2 (1) " "			"*" " "				
##	3 (1) " "		"*"	"*" " "				
##	4 (1) " "	"*"	"*"	"*" " "				
##	5 (1) " "	"*"	"*"	"*" " "				
##	6 (1) "*"	"*"	"*"	"*" " "				
##	7 (1) "*"	"*"	"*"	"*" "*"				
##	8 (1) "*"	"*"	"*"	"*" "*"				
	Nate Wells (Math 243:	Stat Learnin	g)	Subset Selection				

Other Selection Metrics

```
The summary function can return selection metrics for each model.
```

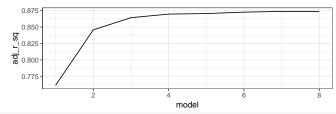
```
adj_r_sq<-summary(best_subset)$adjr2
rss<-summary(best_subset)$rss
cp<-summary(best_subset)$cp</pre>
```

```
d<-data.frame(model = 1:8, adj_r_sq, rss, cp )
d</pre>
```

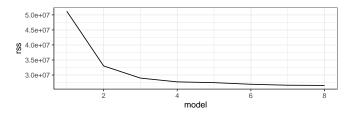
##		model	adj_r_sq	rss	cp
##	1	1	0.7613734	51211963	294.805584
##	2	2	0.8457078	33012815	75.124367
##	3	3	0.8642104	28965893	27.829395
##	4	4	0.8697020	27709979	14.531285
##	5	5	0.8704945	27457472	13.455534
##	6	6	0.8726606	26915647	8.855638
##	7	7	0.8737834	26596486	6.967990
##	8	8	0.8737208	26527820	8.131576

Plotting

We can use ggplot2 to visualize selection metric as a function of variable number ggplot(d, aes(x = model, y = adj_r_sq))+geom_line()+theme_bw()



ggplot(d, aes(x = model, y = rss))+geom_line()+theme_bw()



Forward selection is a computationally efficient alternative to best subset

• To perform forward selection, create the best 1 variable model. Then create p-1 new 2 variable models by adding each other predictor one-at-a-time to the existing 1-variable model. Repeat for 3 variables and so on.

- To perform forward selection, create the best 1 variable model. Then create p-1 new 2 variable models by adding each other predictor one-at-a-time to the existing 1-variable model. Repeat for 3 variables and so on.
- Compared to Best Subset, forward selection computation time grows polynomially in *p*: Num. Models = $1 + \frac{p(p+1)}{2}$

- To perform forward selection, create the best 1 variable model. Then create p-1 new 2 variable models by adding each other predictor one-at-a-time to the existing 1-variable model. Repeat for 3 variables and so on.
- Compared to Best Subset, forward selection computation time grows polynomially in *p*: Num. Models = $1 + \frac{p(p+1)}{2}$
- Forward selection tends to favor parsimonous models

- To perform forward selection, create the best 1 variable model. Then create p-1 new 2 variable models by adding each other predictor one-at-a-time to the existing 1-variable model. Repeat for 3 variables and so on.
- Compared to Best Subset, forward selection computation time grows polynomially in *p*: Num. Models = $1 + \frac{p(p+1)}{2}$
- Forward selection tends to favor parsimonous models
- Downsides?

- To perform forward selection, create the best 1 variable model. Then create p-1 new 2 variable models by adding each other predictor one-at-a-time to the existing 1-variable model. Repeat for 3 variables and so on.
- Compared to Best Subset, forward selection computation time grows polynomially in *p*: Num. Models = $1 + \frac{p(p+1)}{2}$
- Forward selection tends to favor parsimonous models
- Downsides?
 - Not guaranteed to find the best model (or even something close to the best model)

- To perform forward selection, create the best 1 variable model. Then create p-1 new 2 variable models by adding each other predictor one-at-a-time to the existing 1-variable model. Repeat for 3 variables and so on.
- Compared to Best Subset, forward selection computation time grows polynomially in *p*: Num. Models = $1 + \frac{p(p+1)}{2}$
- Forward selection tends to favor parsimonous models
- Downsides?
 - Not guaranteed to find the best model (or even something close to the best model)
 - Early predictors may become redundant

- To perform forward selection, create the best 1 variable model. Then create p-1 new 2 variable models by adding each other predictor one-at-a-time to the existing 1-variable model. Repeat for 3 variables and so on.
- Compared to Best Subset, forward selection computation time grows polynomially in *p*: Num. Models = $1 + \frac{p(p+1)}{2}$
- Forward selection tends to favor parsimonous models
- Downsides?
 - Not guaranteed to find the best model (or even something close to the best model)
 - Early predictors may become redundant
 - Can be unstable

Backward Elimination

Backward Elimination is another computationally efficient alternative to best subset

Backward Elimination

Backward Elimination is another computationally efficient alternative to best subset

• To perform backward selection, begin with full model. Then create p-1 new p-1 variable models by removing one-at-a-time each other predictor from the existing *p*-variable model. Repeat for p-2 variables and so on.

Backward Elimination

Backward Elimination is another computationally efficient alternative to best subset

- To perform backward selection, begin with full model. Then create p-1 new p-1 variable models by removing one-at-a-time each other predictor from the existing *p*-variable model. Repeat for p-2 variables and so on.
- Compared to Best Subset, backward elimination computation time grows polynomially in *p*: Num. Models = $1 + \frac{p(p+1)}{2}$

- To perform backward selection, begin with full model. Then create p-1 new p-1 variable models by removing one-at-a-time each other predictor from the existing *p*-variable model. Repeat for p-2 variables and so on.
- Compared to Best Subset, backward elimination computation time grows polynomially in p: Num. Models = 1 + \frac{p(p+1)}{2}
- Backward elimination tends to favor in-depth models

- To perform backward selection, begin with full model. Then create p-1 new p-1 variable models by removing one-at-a-time each other predictor from the existing *p*-variable model. Repeat for p-2 variables and so on.
- Compared to Best Subset, backward elimination computation time grows polynomially in p: Num. Models = 1 + \frac{\rho(p+1)}{2}
- Backward elimination tends to favor in-depth models
- Downsides?

- To perform backward selection, begin with full model. Then create p-1 new p-1 variable models by removing one-at-a-time each other predictor from the existing *p*-variable model. Repeat for p-2 variables and so on.
- Compared to Best Subset, backward elimination computation time grows polynomially in p: Num. Models = 1 + \frac{\rho(p+1)}{2}
- Backward elimination tends to favor in-depth models
- Downsides?
 - Not guaranteed to find the best model (or even something close to the best model)

- To perform backward selection, begin with full model. Then create p-1 new p-1 variable models by removing one-at-a-time each other predictor from the existing *p*-variable model. Repeat for p-2 variables and so on.
- Compared to Best Subset, backward elimination computation time grows polynomially in p: Num. Models = 1 + \frac{p(p+1)}{2}
- Backward elimination tends to favor in-depth models
- Downsides?
 - Not guaranteed to find the best model (or even something close to the best model)
 - Requires fewer predictors than observations

- To perform backward selection, begin with full model. Then create p-1 new p-1 variable models by removing one-at-a-time each other predictor from the existing *p*-variable model. Repeat for p-2 variables and so on.
- Compared to Best Subset, backward elimination computation time grows polynomially in p: Num. Models = 1 + \frac{p(p+1)}{2}
- Backward elimination tends to favor in-depth models
- Downsides?
 - Not guaranteed to find the best model (or even something close to the best model)
 - Requires fewer predictors than observations
 - Susceptible to multicollinearity

- To perform backward selection, begin with full model. Then create p-1 new p-1 variable models by removing one-at-a-time each other predictor from the existing *p*-variable model. Repeat for p-2 variables and so on.
- Compared to Best Subset, backward elimination computation time grows polynomially in p: Num. Models = 1 + \frac{p(p+1)}{2}
- Backward elimination tends to favor in-depth models
- Downsides?
 - Not guaranteed to find the best model (or even something close to the best model)
 - Requires fewer predictors than observations
 - Susceptible to multicollinearity
 - Can be unstable

Forward/Backward Selection in R

Summary of Forward Selection

summary(forward_select)

##	Subset selection object			
##	Call: regsubsets.formula(body_mass_g ~ ., data = penguins, nvmax = 8,			
##	method = "forward")			
##	9 Variables (and intercept)			
##	Forced in Forced out			
##	speciesChinstrap	FALSE	FALSE	
##	speciesGentoo	FALSE	FALSE	
##	islandDream	FALSE	FALSE	
##	islandTorgersen	FALSE	FALSE	
##	bill_length_mm	FALSE	FALSE	
##	bill_depth_mm	FALSE	FALSE	
##	flipper_length_mm	FALSE	FALSE	
	sexmale	FALSE	FALSE	
	year	FALSE	FALSE	
	1 subsets of each size up to 8			
##	Selection Algorithm: forward			
##				ream islandTorgersen
##	,			
##	- (-)			
##	,	"*"		
##	,	"*"		
##	/	"*"		
##	/	"*"		
##		"*"		
##	/	"*"		"*"
##				ength_mm sexmale year
##	,		"*"	
##	- (-)		"*"	******
##	,		"*"	******
##	/	"*"	"*"	*******
##	/	"*"	"*"	
##	/	"*"	"*"	*
##		"*"	"*"	"*" "*"
##	/	"*"	"*"	"*" "*" Subset Selection
Nate Wells (Math 243: Stat Learning) Subset Selection				

Summary of Backward Elimination

summary(backward elim)

Subset selection object ## Call: regsubsets.formula(body_mass_g ~ ., data = penguins, nvmax = 8, ## method = "backward") ## 9 Variables (and intercept) ## Forced in Forced out ## speciesChinstrap FALSE FALSE ## speciesGentoo FALSE FALSE FALSE FALSE ## islandDream ## islandTorgersen FALSE FALSE ## bill_length_mm FALSE FALSE ## bill_depth_mm FALSE FALSE FALSE FALSE ## flipper_length_mm ## sexmale FALSE FALSE ## vear FALSE FALSE ## 1 subsets of each size up to 8 ## Selection Algorithm: backward ## speciesChinstrap speciesGentoo islandDream islandTorgersen "*" ## 1 (1)"" ## 2 (1) . . ال ب ال ## 3 (1) . . "*" (1) . . "*" . . ## 4 . . ## 5 (1) "*" ال ب ال . . (1) "*" ## 6 "*" . . (1) "*" ## 7 ال ب ال "*" . . ## 8 (1) "*" "*" ## bill length mm bill depth mm flipper length mm sexmale vear . . ## 1 (1) " " ## 2 (1) "*" ## 3 (1)"" "*" 11 - 11 . . ## 4 (1) . . "*" "*" "*" ## 5 (1)"" 11 **a** 11 11 **a** 11 "*" . . 11 yr 11 11 **a** 11 11 yr 11 . . ## 6 (1) "*" ## 7 (1) "*" 11 yr 11 11 **a** 11 11 yr 11 11 <u>a</u> 11 ## 8 (1) "*" 11 yr 11 11 **a** 11 11 yr 11 11 <u>a</u> 11

Section 2

Penalized Regression

Suppose we wish to build a linear model for Y using predictors X_1, \ldots, X_p using n observations.

- Suppose we wish to build a linear model for Y using predictors X_1, \ldots, X_p using n observations.
- Generally, under what circumstances will the **full model** perform well, compared to other subset models? (*Think about the Bias-Variance Tradeoff*)

- Suppose we wish to build a linear model for Y using predictors X_1, \ldots, X_p using n observations.
- Generally, under what circumstances will the **full model** perform well, compared to other subset models? (*Think about the Bias-Variance Tradeoff*)
 - If irreducible error is small ($\epsilon \sim N(0,\sigma^2)$ with $\sigma pprox 0$)

- Suppose we wish to build a linear model for Y using predictors X_1, \ldots, X_p using n observations.
- Generally, under what circumstances will the **full model** perform well, compared to other subset models? (*Think about the Bias-Variance Tradeoff*)
 - If irreducible error is small ($\epsilon \sim N(0, \sigma^2)$ with $\sigma \approx 0$)
 - If model variance is high and model bias is low.

- Suppose we wish to build a linear model for Y using predictors X_1, \ldots, X_p using n observations.
- Generally, under what circumstances will the **full model** perform well, compared to other subset models? (*Think about the Bias-Variance Tradeoff*)
 - If irreducible error is small ($\epsilon \sim N(0, \sigma^2)$ with $\sigma \approx 0$)
 - If model variance is high and model bias is low.

Suppose $\hat{y} = 10 + 0.01x_2 + 1000x_2$ is the best fitting linear model using X_1 and X_2 .

Suppose we wish to build a linear model for Y using predictors X_1, \ldots, X_p using n observations.

Generally, under what circumstances will the **full model** perform well, compared to other subset models? (*Think about the Bias-Variance Tradeoff*)

- If irreducible error is small ($\epsilon \sim N(0,\sigma^2)$ with $\sigma pprox 0$)
- If model variance is high and model bias is low.

Suppose $\hat{y} = 10 + 0.01x_2 + 1000x_2$ is the best fitting linear model using X_1 and X_2 . How might the bias and variance of the following model compare?

 $\hat{y} = 10 + 0.01x_2 + 500x_2$

Suppose we wish to build a linear model for Y using predictors X_1, \ldots, X_p using n observations.

Generally, under what circumstances will the **full model** perform well, compared to other subset models? (*Think about the Bias-Variance Tradeoff*)

- If irreducible error is small ($\epsilon \sim N(0,\sigma^2)$ with $\sigma pprox 0$)
- If model variance is high and model bias is low.

Suppose $\hat{y} = 10 + 0.01x_2 + 1000x_2$ is the best fitting linear model using X_1 and X_2 . How might the bias and variance of the following model compare?

$$\hat{y} = 10 + 0.01x_2 + 500x_2$$

When might this new model have lower MSE than the original model?

Suppose $\hat{y} = 10 + 0.01x_1 + 1000x_2$ is again the best fitting linear model using X_1 and X_2 .

Suppose $\hat{y} = 10 + 0.01x_1 + 1000x_2$ is again the best fitting linear model using X_1 and X_2 .

• Are we justified in saying that X_2 is a more important predictor than X_1 ?

Suppose $\hat{y} = 10 + 0.01x_1 + 1000x_2$ is again the best fitting linear model using X_1 and X_2 .

• Are we justified in saying that X_2 is a more important predictor than X_1 ?

Suppose we first standardize X_1 and X_2 by subtacting off their means and dividing by their standard deviations:

$$Z_1 = rac{X_1 - \mu_1}{\sigma_1}$$
 $Z_2 = rac{X_2 - \mu_2}{\sigma_2}$

Suppose $\hat{y} = 10 + 0.01x_1 + 1000x_2$ is again the best fitting linear model using X_1 and X_2 .

• Are we justified in saying that X_2 is a more important predictor than X_1 ?

Suppose we first standardize X_1 and X_2 by subtacting off their means and dividing by their standard deviations:

$$Z_1 = rac{X_1 - \mu_1}{\sigma_1}$$
 $Z_2 = rac{X_2 - \mu_2}{\sigma_2}$

• If we build a model and find $\hat{y} = 10 + 0.01z_1 + 1000z_2$, where Z_1 and Z_2 are standardized, are we now justified in saying that Z_2 is more important than Z_1 ?

Suppose $\hat{y} = 10 + 0.01x_1 + 1000x_2$ is again the best fitting linear model using X_1 and X_2 .

• Are we justified in saying that X_2 is a more important predictor than X_1 ?

Suppose we first standardize X_1 and X_2 by subtacting off their means and dividing by their standard deviations:

$$Z_1 = rac{X_1 - \mu_1}{\sigma_1}$$
 $Z_2 = rac{X_2 - \mu_2}{\sigma_2}$

- If we build a model and find $\hat{y} = 10 + 0.01z_1 + 1000z_2$, where Z_1 and Z_2 are standardized, are we now justified in saying that Z_2 is more important than Z_1 ?
 - How might the variance and bias of the following model compare to the standarized model?

$$\hat{y} = 10 + 0.02z_2 + 500z_1$$

Recall that least squares regression estimates $\hat{\beta}_0, \hat{\beta}_1, \ldots, \hat{\beta}_p$ for

$$\hat{y} = \beta_0 + \beta_1 X_1 + \dots + \beta_p X_p + \epsilon$$

are obtained by finding the values of β that minimize

$$RSS = \sum_{i=1}^{n} (y_i - \hat{y}_i)^2 = \sum_{i=1}^{n} \left(y_i - \beta_0 - \sum_{j=1}^{p} \beta_j x_{ij} \right)^2$$

Recall that least squares regression estimates $\hat{\beta}_0, \hat{\beta}_1, \ldots, \hat{\beta}_p$ for

$$\hat{y} = \beta_0 + \beta_1 X_1 + \dots + \beta_p X_p + \epsilon$$

are obtained by finding the values of β that minimize

$$RSS = \sum_{i=1}^{n} (y_i - \hat{y}_i)^2 = \sum_{i=1}^{n} \left(y_i - \beta_0 - \sum_{j=1}^{p} \beta_j x_{ij} \right)^2$$

To perform Ridge Regression, we instead find coefficients β that minimize

$$\operatorname{RSS} + \lambda \sum_{i=1}^{p} \beta_{i}^{2}$$
 where $\lambda \geq 0$ is tuning parameter

Recall that least squares regression estimates $\hat{\beta}_0, \hat{\beta}_1, \dots, \hat{\beta}_p$ for

$$\hat{y} = \beta_0 + \beta_1 X_1 + \dots + \beta_p X_p + \epsilon$$

are obtained by finding the values of β that minimize

$$RSS = \sum_{i=1}^{n} (y_i - \hat{y}_i)^2 = \sum_{i=1}^{n} \left(y_i - \beta_0 - \sum_{j=1}^{p} \beta_j x_{ij} \right)^2$$

To perform Ridge Regression, we instead find coefficients β that minimize

$$\operatorname{RSS} + \lambda \sum_{i=1}^{p} \beta_{i}^{2}$$
 where $\lambda \geq 0$ is tuning parameter

Why?

Recall that least squares regression estimates $\hat{\beta}_0, \hat{\beta}_1, \ldots, \hat{\beta}_{\it P}$ for

$$\hat{y} = \beta_0 + \beta_1 X_1 + \dots + \beta_p X_p + \epsilon$$

are obtained by finding the values of β that minimize

$$RSS = \sum_{i=1}^{n} (y_i - \hat{y}_i)^2 = \sum_{i=1}^{n} \left(y_i - \beta_0 - \sum_{j=1}^{p} \beta_j x_{ij} \right)^2$$

To perform Ridge Regression, we instead find coefficients β that minimize

$$\operatorname{RSS} + \lambda \sum_{i=1}^{p} \beta_{i}^{2}$$
 where $\lambda \geq 0$ is tuning parameter

Why?

• The term $\lambda \sum_{i=1}^{p} \beta_i^2$ is the **shrinkage penalty**, and is small when the β are small.

Recall that least squares regression estimates $\hat{\beta}_0, \hat{\beta}_1, \ldots, \hat{\beta}_{\it P}$ for

$$\hat{y} = \beta_0 + \beta_1 X_1 + \dots + \beta_p X_p + \epsilon$$

are obtained by finding the values of β that minimize

$$RSS = \sum_{i=1}^{n} (y_i - \hat{y}_i)^2 = \sum_{i=1}^{n} \left(y_i - \beta_0 - \sum_{j=1}^{p} \beta_j x_{ij} \right)^2$$

To perform Ridge Regression, we instead find coefficients β that minimize

$$\operatorname{RSS} + \lambda \sum_{i=1}^{p} \beta_{i}^{2}$$
 where $\lambda \geq 0$ is tuning parameter

Why?

- The term $\lambda \sum_{i=1}^{p} \beta_i^2$ is the **shrinkage penalty**, and is small when the β are small.
- With a shrinkage penalty, the algorithm prefers models with lower coefficients.

Recall that least squares regression estimates $\hat{\beta}_0, \hat{\beta}_1, \ldots, \hat{\beta}_{\it P}$ for

$$\hat{y} = \beta_0 + \beta_1 X_1 + \dots + \beta_p X_p + \epsilon$$

are obtained by finding the values of β that minimize

$$RSS = \sum_{i=1}^{n} (y_i - \hat{y}_i)^2 = \sum_{i=1}^{n} \left(y_i - \beta_0 - \sum_{j=1}^{p} \beta_j x_{ij} \right)^2$$

To perform Ridge Regression, we instead find coefficients β that minimize

$$\operatorname{RSS} + \lambda \sum_{i=1}^{p} \beta_{i}^{2}$$
 where $\lambda \geq 0$ is tuning parameter

Why?

- The term $\lambda \sum_{i=1}^{p} \beta_i^2$ is the **shrinkage penalty**, and is small when the β are small.
- With a shrinkage penalty, the algorithm prefers models with lower coefficients.
- This tends to reduce variance, at the cost of increased bias.

Goal: Find β which minimize $\text{RSS} + \lambda \sum_{i=1}^{p} \beta_p^2$

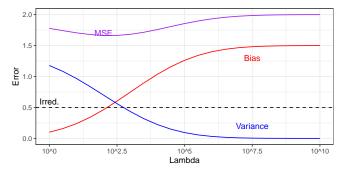
Goal: Find β which minimize $\text{RSS} + \lambda \sum_{i=1}^{p} \beta_p^2$

What will happen to β_0 as $\lambda \to \infty$? As $\lambda \to 0$?

- Goal: Find β which minimize RSS + $\lambda \sum_{i=1}^{p} \beta_{p}^{2}$
- What will happen to β_0 as $\lambda \to \infty$? As $\lambda \to 0$?
- What happens to MSE as $\lambda \rightarrow 0$ or $\lambda \rightarrow 1$?

Goal: Find β which minimize $\text{RSS} + \lambda \sum_{i=1}^{p} \beta_{p}^{2}$ What will happen to β_{0} as $\lambda \to \infty$? As $\lambda \to 0$?

What happens to MSE as $\lambda \rightarrow 0$ or $\lambda \rightarrow 1$?



Suppose X_1 and X_2 are non-standardized predictors and the best fitting linear model is

 $\hat{y} = 10 + 0.01x_1 + 1000x_2$

Suppose X_1 and X_2 are non-standardized predictors and the best fitting linear model is

 $\hat{y} = 10 + 0.01x_1 + 1000x_2$

Suppose X_1 and X_2 are non-standardized predictors and the best fitting linear model is

 $\hat{y} = 10 + 0.01x_1 + 1000x_2$

- Recall the shrinkage penalty is $\lambda \sum_{i=1}^{2} \beta_i^2 = \lambda (1000^2 + 0.01^2)$
- Since $\beta_2 = 1000$ is much larger than $\beta_1 = 0.01$, ridge regression will prioritize reducing β_2 over β_1 .

Suppose X_1 and X_2 are non-standardized predictors and the best fitting linear model is

 $\hat{y} = 10 + 0.01x_1 + 1000x_2$

- Recall the shrinkage penalty is $\lambda \sum_{i=1}^{2} \beta_i^2 = \lambda (1000^2 + 0.01^2)$
- Since $\beta_2 = 1000$ is much larger than $\beta_1 = 0.01$, ridge regression will prioritize reducing β_2 over β_1 .
- Will this actually produce good MSE for a fixed λ?

Suppose X_1 and X_2 are non-standardized predictors and the best fitting linear model is

 $\hat{y} = 10 + 0.01x_1 + 1000x_2$

- Recall the shrinkage penalty is $\lambda \sum_{i=1}^{2} \beta_i^2 = \lambda (1000^2 + 0.01^2)$
- Since $\beta_2 = 1000$ is much larger than $\beta_1 = 0.01$, ridge regression will prioritize reducing β_2 over β_1 .
- Will this actually produce good MSE for a fixed λ ?
 - Only if standard deviation of x₂ is much, much larger than that of x₁

Suppose X_1 and X_2 are non-standardized predictors and the best fitting linear model is

 $\hat{y} = 10 + 0.01x_1 + 1000x_2$

What type of models will ridge regression prefer (in terms of β_1 and β_2)?

- Recall the shrinkage penalty is $\lambda \sum_{i=1}^{2} \beta_i^2 = \lambda (1000^2 + 0.01^2)$
- Since $\beta_2 = 1000$ is much larger than $\beta_1 = 0.01$, ridge regression will prioritize reducing β_2 over β_1 .
- Will this actually produce good MSE for a fixed λ ?
 - Only if standard deviation of x_2 is much, much larger than that of x_1

Ridge regression is most efficient if predictors are standardarized first.

Suppose X_1 and X_2 are non-standardized predictors and the best fitting linear model is

 $\hat{y} = 10 + 0.01x_1 + 1000x_2$

What type of models will ridge regression prefer (in terms of β_1 and β_2)?

- Recall the shrinkage penalty is $\lambda \sum_{i=1}^{2} \beta_i^2 = \lambda (1000^2 + 0.01^2)$
- Since $\beta_2 = 1000$ is much larger than $\beta_1 = 0.01$, ridge regression will prioritize reducing β_2 over β_1 .
- Will this actually produce good MSE for a fixed λ ?
 - Only if standard deviation of x_2 is much, much larger than that of x_1

Ridge regression is most efficient if predictors are standardarized first.

$$\mathsf{z}_{ij} = rac{\mathsf{x}_{ij} - ar{\mathsf{x}}_j}{\hat{\sigma}_j}$$

Where x_{ij} is the *i*th observation of the *j*th predictor, \bar{x}_j is the sample mean of the *j*th predictor, and $\hat{\sigma}_j$ is the sample st. dev. of the *j*th predictor.